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SHORT COMMUNICATION

Two-Stage Conditional Repetitive Group Sampling Plan

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ABSTRACT

This paper enlarges the concept of conditional repetitive group sampling plan. The operating characteristic and average sample number functions of the generalised plan have been derived by graphical evaluation and review technique. Lastly, Poisson unity values have been tabulated to facilitate the operation and construction of the plan. The plan that provides a quick discrimination of good and bad quality lots, could be used for both the process control and goods acceptance or rejection in various stages of production in defence organisations.

Keywords: Conditional repetitive group sampling plan, Poisson unity value, graphical evaluation and review technique, repetitive group sampling plan

1. INTRODUCTION

The concept of repetitive group sampling (RGS) plan was introduced by Sherman¹ in which acceptance and rejection of the lot was based on the repeated sample results of the same lot. The detailed procedures and tables for construction and selection of RGS plan have been given by Soundararajan and Ramaswamy², and Singh³, et. al. Later on, Shankar and Mohapatra⁴ developed conditional RGS plan as an extension of the classical RGS plan. Mohapatra⁵ compared the conditional RGS plan with the RGS plan and observed that the conditional RGS plan is better in sample size efficiency than the RGS plan. The purpose of present investigation is two-fold. Firstly, following Stephens and Dodge⁶, the proposed plan uses different sample sizes in the normal and the tightened phases of inspection. Secondly, the dynamic characteristics of the proposed plan have been modelled and analysed through graphical evaluation and review technique⁷⁻⁹ (GERT), which has been used by

several authors to study quality control systems. A brief account of such studies has been given by Shankar¹⁰. The formula for operating characteristic and average sample number (ASN) functions of the plan has been derived by applying Mason's¹¹ rule on the GERT network representation of the inspection system. Lastly, Poisson unity values have been tabulated to facilitate the operation and construction of the plan.

Military standard plans are commonly used in defence establishments by the military, as a consumer. In application of the MIL-STD-105D system, it is intended that a switch to tightened inspection, in case of poor quality, provides a psychological and economic incentive for the producer to improve the level of the product quality submitted. The proposed two-stage plan based on normal and tightened inspection may be useful as an addition to MIL-STD-105D system for quick discrimination of good and bad quality lots. The present study may also be used when output from a continuous process is inspected and/or sampled continuously to determine the quality of goods. Thus, the proposed procedure could be used for both the process control and goods acceptance or rejection in various stages of production in defence organisations.

2. TWO-STAGE CONDITIONAL REPETITIVE GROUP SAMPLING PLAN

Following the notations and concepts similar to those of Sherman, and Shankar and Mohapatra, the proposed two-stage conditional RGS plan is carried out through the following steps:

- (a) Draw a random sample of size n_1 from the lot for normal inspection and determine the number of defectives (d) found therein.
- (b) Accept the lot, if $d \le c_1$

Reject the lot, if $d > c_2$

(c) If
$$c_1 < d \le c_2$$
, then

- Switch to tightened inspection by taking a sample of size $n_2 (> n_1)$ from the lot and determine the number of defectives (d) found therein.
- Accept the lot, if $d \leq c_1$
- Reject the lot, if $d > c_{2}$.
- If $c_1 < d \le c_2$, then repeat the above steps provided previous *i* lots are accepted under normal inspection.
- Otherwise reject the lot.

Here, it may be noted that a lot with number of defectives $\leq c_1$ is accepted at both the normal and tightened inspection states. Furthermore, the process is automatically switched to normal inspection after acceptance/rejection of the current lot.

The proposed plan is characterised by five parameters, namely n_1, n_2, c_1, c_2 and *i*. For i = 0, the resulting plan is two-phase inspection RGS plan due to Shankar¹². Moreover, for i = 0 and $k = n_2/n_1 = 1$ (i.e. $n_1 = n_2$), one has RGS plan due to Sherman¹.

3. GRAPHICAL EVALUATION & REVIEW TECHNIQUE - ANALYSIS OF PLAN

The possible stages of the inspection system can be defined as follows:

- S_0 Initial stage of the plan
- S₁ Stage in which tightened inspection is performed for the current lot.
- $S_2(j)$ Stage in which j^{th} (j = 1, 2, 3, ..., i) preceding lot is accepted under normal inspection.
- S_{\star} Stage in which current lot is accepted.
- S_{p} Stage in which current lot is rejected.

The above stages enable one to construct GERT network representation of the inspection system as shown in Fig. (1). The moment generating function of sample size n is:

$$M_n(\theta) = \exp n(\theta)$$

because *n* is constant. Now, applying Mason's¹¹ rule on the representation in Fig. (1), the *W*-function⁷⁻⁹ from the initial node S_0 to the terminal nodes S_A and S_R , respectively is given by

$$W_{A}(\theta) = \left[P_{a_{1}} e^{n_{1}\theta} \left\{ 1 - P_{c_{2}} e^{n_{2}\theta} \left(P_{a_{1}} e^{n_{1}\theta} \right)^{i} \right\} + P_{c_{1}} P_{a_{2}} e^{(n_{1}+n_{2})\theta} \right] / H$$
$$W_{R}(\theta) = \left[P_{r_{1}} e^{n_{1}\theta} \left\{ 1 - P_{c_{2}} e^{n_{2}\theta} \left(P_{a_{1}} e^{n_{1}\theta} \right)^{i} \right\} + P_{c_{1}} P_{a_{2}} e^{(n_{1}+n_{2})\theta} + G \right] / H^{(1)}$$

where

$$\begin{split} H &= 1 - P_{c_2} e^{n_1 \theta} \left(P_{a_1} e^{n_1 \theta} \right)^i \\ J &= P_{c_1} P_{c_2} (1 - P_{a_1}) e^{(2n_1 + n_2) \theta} \left\{ 1 - \left(P_{a_1} e^{n_1 \theta} \right)^i / (1 - P_{a_1} e^{n_1 \theta}) \right\} \\ P_{a_1} &= L(p, n_1, c_1), \qquad P_{a_2} = L(p, n_2, c_1) \\ P_{r_1} &= 1 - L(p, n_1, c_2), \quad P_{r_2} = 1 - L(p, n_2, c_2) \\ P_{c_1} &= 1 - (P_{a_1} + P_{r_1}), \qquad P_{c_2} = 1 - (P_{a_2} + P_{r_2}) \end{split}$$

and

L(p, n, c) is the probability of getting c or less defectives with product quality p and sample size n. GAURI SHANKAR et al: TWO-STAGE CONDITIONAL REPETITIVE GROUP SAMPLING PLAN



Figure 1. GERT network of two-stage conditional RGS plan

Now, from the definition of W-function, one gets:

$$M_{\mathcal{A}}(\theta) = W_{\mathcal{A}}(\theta) / W_{\mathcal{A}}(0)$$

$$P_{A} = \left[P_{a_{1}} \left(1 - P_{c_{2}} P_{a_{1}}^{i} \right) + P_{c_{1}} P_{a_{2}} \right] / \left(1 - P_{c_{2}} P_{a_{1}}^{i} \right)$$

$$P_{R} = \left[P_{r_{1}} \left(1 - P_{c_{2}} P_{a_{1}}^{i} \right) + P_{c_{1}} P_{r_{2}} + P_{c_{1}} P_{c_{2}} \left(1 - P_{a_{1}}^{i} \right) \right] / \left(1 - P_{c_{2}} P_{a_{1}}^{i} \right)$$
(2)

Thus, $P_R = 1 - P_A$, where P_A and P_R stand for the probability of acceptance and rejection of the lot, respectively.

For further characterisation of the plan, ASN function is defined as follows:

$$ASN = P_{A} \left[\frac{d}{d\theta} M_{A}(\theta) \right]_{\theta=0} + P_{R} \left[\frac{d}{d\theta} M_{R}(\theta) \right]_{\theta=0}$$
$$= \left[n_{1} \left\{ 1 - P_{c_{2}} P_{a_{1}}^{i} + P_{c_{1}} P_{c_{2}} \left(1 - P_{a_{1}} \right)^{-1} \right]_{\theta=0} (3)$$
$$(1 - P_{a_{1}}^{i}) + n_{2} P_{c_{1}} \right] / (1 - P_{c_{2}} P_{a_{1}}^{i})$$

where

$$M_{p}(\theta) = W_{p}(\theta) / W_{p}(\theta)$$

Here, it may be observed that for i = 0, one has:

$$P_{A} = P_{a_{1}} + P_{c_{1}} P_{a_{2}} / (1 - P_{c_{2}})$$

and

$$ASN = n_1 + n_2 P_{c_1} / (1 - P_{c_2})$$
(4)

These results agree with Shankar¹². Furthermore, for k = 1 i.e. $n_1 = n_2$

$$P_{a_1} = P_{a_2}, P_{a_1} = P_{a_2}, \text{ and } P_{a_1} = P_{a_2}$$

Therefore

$$P_{A} = P_{a_{1}} / (1 - P_{c_{1}})$$

and

Table 1. Values of $\mathbf{R} = p_2/p_1$ and n_1p_2 for two-stage conditional RGS plan, where $\alpha = 0.05$, $\beta = 0.10$, $c_1 = 0$, $c_2 = 2$

	k = 1.00		k = 1.25		k = 1.50		<i>k</i> = 1.75		k = 2.00	
i	R	n _µ p ₂	R	$n_{\mu}p_{2}$	Ŕ	n_1p_2	R	n_1p_2	R	n_1p_2
1	06.894	2.6697	07.109	2.5175	07.402	2.4365	07.776	2.3829	08.072	2.3496
2	08.014	2.6615	08.276	2.5218	08.549	2.4307	08.885	2.3800	09.241	2.3478
3	08.653	2.6605	09.033	2.5216	09.296	2.4281	09.673	2.3783	10.065	2.3464
4	09.158	2.6601	09.544	2.5216	09.906	2.4260	10.331	2.3769	10.745	2.3453
5	09.518	2.6599	09.918	2.5216	10.327	2.4342	10.727	2.3760	11.153	2.3446
6	09.784	2.6597	10.201	2.5217	10.626	2.4343	11.036	2.3753	11.388	2.3442
7	09.987	2.6595	10.422	2.5217	10.862	2.4343	11.282	2.3748	11.718	2.3436
8	10.143	2 6594	10.596	2.5217	11.051	2.4343	11.370	2.3746	12.017	2.3431
9	10.265	2.6593	10.734	2.5217	11.203	2.4344	11.577	2.3741	12.248	2.3426
10	10.360	2.6593	10.845	2.5217	11.327	2.4344	11.868	2.3735	12.249	2.3426

$$ASN = n_1 / (1 - P_c)$$

These results coincide with Sherman¹

4. CONSTRUCTION OF TABLES

The expression for the probability of acceptance (P_A) and ASN of the two-stage conditional RGS plan under Poisson model are given by Eqns (2) and (3), respectively with the following expressions:

$$P_{a_1} = \sum_{x=0}^{c_1} (n_1 p)^x \exp(-n_1 p) / x$$
$$P_{a_2} = \sum_{x=0}^{c_1} (k n_1 p)^x \exp(-k n_1 p) / x$$

$$P_{n_1} = 1 - \sum_{x=0}^{c_2} (n_1 p)^x \exp(-n_1 p) / x$$

$$P_{r_2} = 1 - \sum_{x=0}^{c_2} (kn_1p)^x \exp(-kn_1p) / x$$

where

$$k = n_2/n_1$$

x=0

Let
$$P_a(p') = L(p', n, c)$$

Now, for given p_1 , p_2 , α and β , i.e. the two

points $(p_1, 1-\alpha)$ and (p_2, β) on the operating characteristic curve, one may write the following expressions:

$$P_A(p_1) = 1 - \alpha \tag{5}$$

)

and

 $P_A(p_2) = \beta \tag{6}$

For given c_1, c_2, k, α and *i*, Eqn (5) is a function of n_1p_1 only. Therefore, unity values n_1p_1 have been obtained from Eqn (5) by Newton's method of successive approximation. Similarly, the values of n_1p_2 are also obtained from Eqn (6) by Newton's method of successive approximation. The operating ratio $R = n_1p_2/n_1p_1 = p_2/p_1$. Only the extract table showing relevant entries for α = 0.05, β = 0.10 and k = 1.00, 1.25, 1.50, 1.75 and 2.00 are presented.

5. RESULTS & DISCUSSION

Suppose a two-stage conditional RGS plan is desired having 95 per cent probability of acceptance at $p_1 = 0.01$ and 10 per cent probability of acceptance at $p_2 = 0.0855$ with $c_1 = 0$, $c_2 = 2$ and k = 1.5.

The operating ratio, $R = p_2/p_1 = 8.55$, where R corresponds to k = 1.5 (Table 1) providing i = 2.

The sample size is:

GAURI SHANKAR et al: TWO-STAGE CONDITIONAL REPETITIVE GROUP SAMPLING PLAN

		Average sample number (ASN)								
p	P	k = 1	k = 1.25	<i>k</i> = 1.50	k = 1.75	k = 2.00				
$p_1 = 0.01$	0.99	27.1167	26.0582	,25.2366	24.5728	24.0270				
	0.95	58.8096	57.4323	56.4059	55.6001	54.9485				
p*=0.05	0.50	51.5196	51.4251	51.4213	51.4800	51.5887				
$p_2 = 0.10$	0.10	43.7175	44.2184	45.2071	46.6148	48.3789				
	0.05	47.1490	47.8127	49.1597	51.0117	53.2199				

Table 2. The effect of different choices of k on operating characteristic and average sample number

 $n_1 = n_1 p_2 / p_2 = 2.4307/0.0855 = 28.429 \simeq 28$ and

 $n_2 = k \times n_1 = 1.5 \times 28.429 = 42.643 \approx 43$.

Thus, the desired plan consists of $c_1 = 0$, . $c_2 = 2$, $n_1 = 28$, $n_2 = 43$ and i = 2.

Table 2 shows the effect of different choices of k on operating characteristic and ASN for some selected values of p, say p_1 , p_2 and p^* with $c_1 = 0$, $c_2 = 2$ and i = 1, where $P_A(p^*) = 0.5$.

It is observed from this table that for the lots of good quality, the ASN of the plan decreases with increasing k. However, the situation is other way for poor quality lots, i.e. for $p = p_2$. Thus, the proposed plan is suitable to the lots of good quality for its quick disposal.

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